

Lecture 4 Random Vector and Expectation

*Lecturer: Xiangyu Chang**Scribe: Hongyi Xu, Ye Zhang**Edited by: Zhihong Liu*

1 Recall

1.1 CDF

- $F_{X,Y}(x, y) = P(x \leq x, y \leq y)$

1.2 PMF and PDF

- PMF:

- $f_{X,Y}(x, y) = P(x \leq x, y \leq y)$

- PDF:

- $f_{X,Y}(x, y) \geq 0$

- $\int_{\mathbb{R}^2} f_{X,Y}(x, y) dx dy = 1$

- $P((X, Y) \in A) = \int_A f_{X,Y}(x, y) dx dy$

1.3 Marginal Distribution

- $f_X(x) = P(X = x) = \sum_y P(X = x, Y = y) = \sum_y f_{X,Y}(x, y)$

- $f_X(x) = \int_{\mathbb{R}} f_{X,Y}(x, y) dy$

1.4 Independent

- $P(X \in A, Y \in B) = P(X \in A)P(X \in B)$

2 Conditional PMF and PDF

- $f_{Y|X}(y) = \frac{f_{X,Y}(x,y)}{f_X(x)}$ or $f_{X,Y}(x, y) = f_Y(y|x) \cdot f_X(x)$

if independent:

- $f_Y(y|x) = f_Y(y)$

For example,

- $X \sim U[0, 1], Y|X \sim U[x, 1], f_Y?$

Step1:

- $f_x(x) = \begin{cases} 1, & 0 \leq x \leq 1 \\ 0, & \text{otherwise} \end{cases}$
- $f_{Y|X}(y) = \begin{cases} \frac{1}{1-x}, & x \leq y < 1 \\ 0, & \text{otherwise} \end{cases}$

Step2:

- $f_{X,Y}(x, y) = f_Y(y|x) \cdot f_X(x) = \begin{cases} \frac{1}{1-x}, & x \leq y < 1 \\ 0, & \text{otherwise} \end{cases}$

Step3:

- $f_Y(y) = \int_{\mathbb{R}} f_{X,Y}(x, y) dx = \int_0^y \frac{1}{1-x} dx = -\ln(1-y)$

supplement: Linear Model

- $\{(x_i, y_i)\}_{i=1}^n$
- $Y_i = \beta^T x_i + \varepsilon_i$

3 Multivariable

3.1 CDF

- $\vec{x} = x = (x_1, \dots, x_d)^\top$
 $F_Y(y) = P(y_1 \leq x_1, \dots, y_d \leq x_d)$

3.2 PMF and PDF

- $F_Y(y) = P(y_1 = x_1, \dots, y_d = x_d)$

- $\int_{R^d} f_x(x) = 1$
- $P(x \in A) = \int_A f_x(x) dx, A \subseteq R^d$
- $f_{x_1 \dots x_d}(x_1, \dots, x_k) = \int_R \int_R \dots f_x(x) dx_{k+1} \dots dx_d$
- $f_x(x) = \prod_{i=1}^d f_{x_i(x_i)}$ (Independent)

3.3 Multivariate Normal Distribution

Definition 1 *Standard Multivariate Normal Distribution*

$Z = (Z_1^\top, Z_2^\top, \dots, Z_k^\top)$, where $Z_1, \dots, Z_k \sim N(0, 1)$ are independent. The density of Z is

$$\begin{aligned} f(z) &= \prod_{i=1}^k f(z_i) \\ &= \frac{1}{(2\pi)^{k/2}} \exp \left\{ -\frac{1}{2} \sum_{j=1}^k z_j^2 \right\} \\ &= \frac{1}{(2\pi)^{k/2}} \exp \left\{ -\frac{1}{2} z^\top z \right\}. \end{aligned}$$

We written that $Z \sim N(0, I)$, I is the $k \times k$ identity matrix.

Definition 2 *(General) Multivariate Normal Distribution*

a vector X has a multivariate normal distribution $X \sim N(\mu, \Sigma)$, it has density

$$f(x; \mu, \Sigma) = \frac{1}{(2\pi)^{k/2} |\Sigma|^{1/2}} \exp \left\{ -\frac{1}{2} (x - \mu)^\top \Sigma^{-1} (x - \mu) \right\}$$

where $|\Sigma|$ denotes the determinant of Σ , μ is a vector of length k and Σ is the $k \times k$ symmetric, positive definite matrix.

Lemma 1 if $X \sim N(0, I)$, $Z = \mu + \Sigma^{1/2} X \sim N(\mu, \Sigma)$

Proof process:

$$\begin{aligned}
&\because Z = \mu + \Sigma^{1/2}X \\
&\because X = g^{-1}(Z) = \Sigma^{-1/2}(Z - \mu) \\
&\nabla g^{-1}(Z) = \Sigma^{-1/2} \\
&f_z(z) = f_x(g^{-1}(z))|\text{Det}(\nabla g^{-1}(Z))| \\
&= \frac{1}{(2\pi)^{k/2}} \exp\left\{-\frac{1}{2}(\Sigma^{-1/2}(Z - \mu))^\top \Sigma^{-1/2}(Z - \mu)\right\}|\text{Det}(\Sigma^{-1/2})| \\
&= \frac{1}{(2\pi)^{k/2}} \exp\left\{-\frac{1}{2}(\Sigma^{-1/2}(Z - \mu))^\top \Sigma^{-1/2}(Z - \mu)\right\}|\text{Det}(\Sigma)|^{-1/2} \\
&= \frac{1}{(2\pi)^{k/2}|\Sigma|^{1/2}} \exp\left\{-\frac{1}{2}(Z - \mu)^\top \Sigma^{-1}(Z - \mu)\right\}
\end{aligned} \tag{1}$$

Definition 3 $\Sigma^{1/2}$ — the square root of Σ
has the following properties:

- $\Sigma^{1/2}$ is symmetric
- $\Sigma = \Sigma^{1/2}\Sigma^{1/2}$
- $\Sigma^{1/2}\Sigma^{-1/2} = \Sigma^{-1/2}\Sigma^{1/2} = I$, where $\Sigma^{-1/2} = (\Sigma^{1/2})^{-1}$
- $\Sigma^{1/2} = UD^{1/2}U^\top$, $\Sigma^{-1/2} = UD^{-1/2}U^\top$ where $D^{1/2} = \begin{pmatrix} \sqrt{D_{11}} & \cdots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \cdots & \sqrt{D_{kk}} \end{pmatrix}$

3.4 Multinomial Distribution

Considering throwing a coin which has k different faces n times.

$p = (p_1, \dots, p_k)$, p_j : the probability of throwing a coin with face j . ($p_j \geq 0$ and $\sum_{j=1}^k p_j = 1$)
 $X = (X_1, \dots, X_k)$, X_j : the number of times that face j appears. ($n = \sum_{j=1}^k X_j$)

We say that X has a *Multinomial*(n, p) distribution written $X \sim \text{Multinomial}(n, p)$.

The probability function is

$$\mathbb{P}(X_1 = n_1, X_2 = n_2, \dots, X_k = n_k) = \frac{n!}{n_1!n_2!\dots n_k!} p_1^{n_1} p_2^{n_2} \dots p_k^{n_k}$$

Lemma 2 Suppose that $X \sim \text{Multinomial}(n, p)$ where $X = (X_1, \dots, X_k)$ and $p = (p_1, \dots, p_k)$.
The marginal distribution of $X_j \sim B(n, p_j)$.

3.5 Expectation

3.5.1 Mean Value

Definition 4 *The expected value, or mean, or first moment, of X is defined to be*

$$\mathbb{E}(X) = \int x dF(x) = \begin{cases} \sum_x x f(x) & \text{if } X \text{ is discrete} \\ \int x f(x) dx & \text{if } X \text{ is continuous} \end{cases}$$

Example 1 *Flip a fair coin two times. Let X be the number of heads. Then,*

$$\begin{aligned} \mathbb{E}(X) &= \int x dF_X(x) \\ &= \sum_x x f_X(x) \\ &= 0 \times f(0) + 1 \times f(1) + 2 \times f(2) \\ &= 0 \times 1/4 + 1 \times 1/2 + 2 \times 1/4 \\ &= 1 \end{aligned} \tag{2}$$

3.5.2 The Rule of the Lazy Statistician

Definition 5 *(The Rule of the Lazy Statistician) Let $Y = r(X)$. Then*

$$\begin{aligned} \mathbb{E}(Y) &= \mathbb{E}(r(X)) \\ &= \int r(x) dF_X(x) \\ &= \int r(x) f_X(x) dx \end{aligned} \tag{3}$$

Example 2 *Let A be an event where $I_A(x) = 1$ if $x \in A$ and $I_A(x) = 0$ if $x \notin A$. Then*

$$\mathbb{E}(I_A(X)) = 0 \cdot \mathbb{P}(X \notin A) + 1 \cdot \mathbb{P}(x \in A) = \mathbb{P}(X \in A).$$